



ELSEVIER

Available at  
[www.ElsevierMathematics.com](http://www.ElsevierMathematics.com)  
POWERED BY SCIENCE @ DIRECT®

Statistics & Probability Letters 65 (2003) 401–410

**STATISTICS &  
PROBABILITY  
LETTERS**

[www.elsevier.com/locate/stapro](http://www.elsevier.com/locate/stapro)

# Chover-type laws of the iterated logarithm for weighted sums

Liang Peng<sup>a</sup>, Yongcheng Qi<sup>b,\*</sup>

<sup>a</sup>*School of Mathematics, Georgia Institute of Technology, Atlanta,  
GA 30332-0160, USA*

<sup>b</sup>*Department of Mathematics and Statistics, University of Minnesota Duluth, SCC 140,  
1117 University Drive, Duluth, MN 55812, USA*

Received March 2003; accepted August 2003

---

## Abstract

In this paper, a Chover-type law of the iterated logarithm is established for the weighted sums of independent and identically distributed random variables with a distribution in the domain of attraction of a stable law.

© 2003 Elsevier B.V. All rights reserved.

MSC: 60F15; 60G50

Keywords: Almost sure convergence; Chover-type law of the iterated logarithm; Stable law; Weighted sum

---

## 1. Introduction

Let  $\{X, X_n, n \geq 1\}$  be a sequence of independent and identically distributed (i.i.d.) random variables with a distribution function  $F(x) = P(X \leq x)$ , and  $h$  be a function defined over  $[0, 1]$ . A weighted sum is defined as

$$S_n(h) = \sum_{k=1}^n h\left(\frac{k}{n}\right) X_k.$$

In case  $h(x) = 1$  for  $x \in [0, 1]$ ,  $S_n := S_n(h) = S_n(1)$  becomes the partial sum.

When  $h$  belongs to a certain class of continuous functions over  $[0, 1]$ , the laws of the iterated logarithm for the weighted sums  $S_n(h)$  have been studied by many authors, for example, Gaposkin (1965), Tomkins (1976), Lai and Wei (1982), Stadtmüller (1984), Li and Tomkins (1996). A recent

---

\* Corresponding author. Fax: +1-218-726-8399.

E-mail address: [yqi@d.umn.edu](mailto:yqi@d.umn.edu) (Y. Qi).

result by Li and Tomkins (1996) showed that

$$\limsup_{n \rightarrow \infty} \frac{S_n(h)}{(2n \log \log n)^{1/2}} = \left( \int_0^1 h(x)^2 dx \right)^{1/2} \quad \text{a.s.} \quad (1.1)$$

if and only if  $E(X) = 0$  and  $E(X^2) = 1$ . Therefore, if  $E(X^2) = \infty$ , an analogue of (1.1) is no longer valid. However, for the partial sum  $S_n$ , Chover (1966) showed that

$$\limsup_{n \rightarrow \infty} \left( \frac{|S_n|}{n^{1/\alpha}} \right)^{1/\log \log n} = e^{1/\alpha} \quad \text{a.s.} \quad (1.2)$$

when  $X$  has a symmetric stable distribution with exponent  $\alpha \in (0, 2)$ , that is,

$$E(\exp(itX)) = \exp(-|t|^\alpha) \quad \text{for } t \in R. \quad (1.3)$$

This motivated the study of Chover-type law of the iterated logarithm for weighted sums in Chen (2002). One of the results in Chen (2002) is the following: under condition (1.3) and certain constraints for  $h$ ,

$$\limsup_{n \rightarrow \infty} \left( \frac{|S_n(h)|}{n^{1/\alpha}} \right)^{1/\log \log n} = e^{1/\alpha} \quad \text{a.s.} \quad (1.4)$$

Note that Qi and Cheng (1996) extended the Chover-type law of the iterated logarithm for the partial sums to the case where the underlying distribution is in the domain of attraction of a stable distribution (see below for details).

Let  $L_\alpha$  denote a stable distribution with exponent  $\alpha \in (0, 2)$ . Recall that the distribution of  $X$  is said to be in the domain of attraction of  $L_\alpha$  if there exist some constants  $A_n \in R$  and  $B_n > 0$  such that

$$\frac{S_n - A_n}{B_n} \xrightarrow{d} L_\alpha. \quad (1.5)$$

Under (1.5), Qi and Cheng (1996) showed that

$$\limsup_{n \rightarrow \infty} \left( \frac{|S_n - A_n|}{B_n} \right)^{1/\log \log n} = e^{1/\alpha} \quad \text{a.s.}$$

It is easy to verify that in the special case (1.3), (1.5) holds with  $A_n = 0$  and  $B_n = n^{1/\alpha}$ , and that  $L_\alpha$  has the same distribution as that of  $X$ . Based on Chen (2002) and Qi and Cheng (1996), we expect that the Chover-type law of the iterated logarithm for weighted sums holds under the general condition (1.5).

We organize this paper as follows. Our main results are presented in Section 2 and all proofs are postponed till Section 3.

## 2. Main results

Let  $BV[0, 1]$  and  $B[0, 1]$  denote sets of all bounded variation functions and of all bounded functions over  $[0, 1]$ , respectively. As in the introduction, let  $\{X, X_n, n \geq 1\}$  be a sequence of i.i.d. random variables satisfying (1.5).

It is well known that (1.5) holds if and only if

$$1 - F(x) = \frac{C_1(x)l(x)}{x^\alpha} \quad \text{and} \quad F(-x) = \frac{C_2(x)l(x)}{x^\alpha}, \quad \text{for } x > 0, \tag{2.1}$$

where, for  $x > 0$ ,  $C_i(x) \geq 0$ ,  $\lim_{x \rightarrow \infty} C_i(x) = C_i$ ,  $i = 1, 2$ ,  $C_1 + C_2 > 0$ , and  $l(x) \geq 0$  is a slowly varying function, i.e.,

$$\lim_{t \rightarrow \infty} \frac{l(tx)}{l(t)} = 1 \quad \text{for } x > 0.$$

Set  $G(x) = P(|X| > x)$ , define

$$B(x) = \inf \left\{ y : G(y) \leq \frac{1}{x} \right\}, \quad x > 0, \tag{2.2}$$

and write

$$\mu_n(c) = n \int \frac{x}{1+x^2} dF(cx) \quad \text{for } c > 0.$$

From Loève (1977), (2.1) is equivalent to the statement that  $(S_n - \mu_n(B(n)))/B(n)$  converges in distribution to a stable distribution with exponent  $\alpha$ . Moreover, we have  $\lim_{n \rightarrow \infty} B_n/B(n) \in (0, \infty)$  and  $\lim_{n \rightarrow \infty} (A_n - \mu_n(B(n)))/B(n) \in R$  from convergence of types theorem. In particular, if  $\alpha > 1$  then  $E(|X|) < \infty$ , and one can always choose  $B_n = B(n)$  and  $A_n = nE(X)$  in (1.5). If  $\alpha < 1$  then one can set  $A_n = 0$  in (1.5).

Without loss of generalization, we always assume that the condition (1.5) holds with  $B(n)$  defined in (2.2) and some constants  $A_n$  with  $A_0 = 0$ . Set  $\lg_0(x) = x$  and  $\lg_k(x) = \log \max(\lg_{k-1}(x), e)$  for  $k \geq 1$ . Our main theorems are as follows.

**Theorem 2.1.** *If  $h \in BV[0, 1]$  with  $h(x_0) \neq 0$  for some  $x_0 \in (0, 1]$ , and  $h$  is continuous at  $x_0$ , then under condition (1.5), for any integer  $r \geq 1$ ,*

$$\limsup_{n \rightarrow \infty} \left( \frac{|S_n(h) - C_n|}{B(\prod_{i=0}^{r-1} \lg_i(n))} \right)^{1/\lg_{r+1}(n)} = e^{1/\alpha} \quad \text{a.s.}, \tag{2.3}$$

where

$$C_n = \sum_{k=1}^n h\left(\frac{k}{n}\right) (A_k - A_{k-1}).$$

**Remark 1.** If  $\alpha \in (1, 2)$ , one can set  $A_n = nE(X)$  in Theorem 2.1, and (2.3) holds with

$$C_n = E(X) \sum_{k=1}^n f\left(\frac{k}{n}\right) = nE(X) \int_0^1 h(x) dx + O(1)$$

since  $h \in BV[0, 1]$ ; if  $\alpha < 1$ , set  $A_n = 0$  and then (2.3) holds with  $C_n = 0$ .

**Remark 2.** Theorem 2.1 generalizes the result (1.4) in Chen (2002) in two directions: (i) the underlying distributions are allowed from the symmetric stable distributions to those in the domain of attraction of a stable law, and (ii) more accurate results than the law of the iterated logarithm are

provided. As a consequence Chover-type law of the iterated logarithm is obtained by taking  $r = 2$  in Theorem 2.1, i.e.,

$$\limsup_{n \rightarrow \infty} \left( \frac{|S_n(h) - C_n|}{B(n)} \right)^{1/\log \log n} = e^{1/\alpha} \quad \text{a.s.}$$

Next we consider more general weighted sums like Chen (2002) did. More specifically, consider an array of weights  $\{a_{n,k}, 1 \leq k \leq n, n \geq 1\}$  satisfying conditions

(C1)  $\sup_{n \geq 1} (\sum_{k=1}^{n-1} |a_{n,k} - a_{n,k+1}| + |a_{n,n}|) < \infty$ ;

(C2) There exist two increasing sequences of  $\{n(k), k \geq 1\}$  and  $\{m(k), k \geq 1\}$  such that  $\sup_{k \geq 1} (n(k+1) - n(k)) < \infty$  and  $\liminf_{k \rightarrow \infty} |a_{n(k),m(k)}| > 0$ .

Note that (C1) is satisfied if  $\sup_{n \geq 1} \max_{1 \leq k \leq n} |a_{n,k}| < \infty$ .

**Theorem 2.2.** *Assume the conditions in Theorem 2.1 and conditions (C1) and (C2) are true. Then*

$$\limsup_{n \rightarrow \infty} \left( \frac{|\sum_{k=1}^n a_{n,k} X_k - C_n|}{B(\prod_{i=0}^{r-1} \lg_i(n))} \right)^{1/\lg_{r+1}(n)} = e^{1/\alpha} \quad \text{a.s.}$$

for all  $r \geq 1$ , where

$$C_n = \sum_{k=1}^n a_{n,k} (A_k - A_{k-1}).$$

A straightforward application of Theorem 2.2 with  $a_{n,k} = \sum_{i=k}^n h(i/n)$  leads to the following theorem.

**Theorem 2.3.** *Let  $h \in B[0, 1]$  with  $\int_{x_0}^1 h(x) \neq 0$  for some  $x_0 \in (0, 1)$ . Then under condition (1.5)*

$$\limsup_{n \rightarrow \infty} \left( \frac{|\sum_{k=1}^n h(\frac{k}{n}) S_k - C_n|}{B(\prod_{i=0}^{r-1} \lg_i(n))} \right)^{1/\lg_{r+1}(n)} = e^{1/\alpha} \quad \text{a.s.,}$$

for any  $r \geq 1$ , where

$$C_n = \sum_{k=1}^n h\left(\frac{k}{n}\right) A_k.$$

As another application of Theorem 2.2, we get the following Chover-type law of the iterated logarithm for moving sums.

**Theorem 2.4.** *Let  $\{b_n\}$  be a sequence of real numbers satisfying  $0 < \sum_{n=1}^{\infty} |b_n| < \infty$ . For  $n \geq 1$  set  $Y_n = \sum_{k=1}^n b_{n-k+1} X_k$ . Then under condition (1.5)*

$$\limsup_{n \rightarrow \infty} \left( \frac{|\sum_{k=1}^n Y_k - C_n|}{B(\prod_{i=0}^{r-1} \lg_i(n))} \right)^{1/\lg_{r+1}(n)} = e^{1/\alpha} \quad \text{a.s.,}$$

for any  $r \geq 1$ , where

$$C_n = \sum_{k=1}^n b_k(A_k - A_{k-1}).$$

### 3. Proofs

We need some lemmas before we proceed to the proofs of our theorems.

**Lemma 3.1.** *Suppose  $l(x)$  is a slowly varying function at infinity and  $g(x) \geq 1$  is an arbitrary function. Then for any given  $\delta > 0$ , there exists an  $x_0 > 0$  such that*

$$\frac{1}{2}g^{-\delta}(x) < \inf_{x \leq y \leq xg(x)} \frac{l(y)}{l(x)} \leq \sup_{x \leq y \leq xg(x)} \frac{l(y)}{l(x)} < 2g^\delta(x) \quad \text{for all } x > x_0.$$

**Lemma 3.2.** *Assume that  $\{Y_n, n \geq 1\}$  is a sequence of i.i.d. random variables and  $\{a_n\}$  is a sequence of positive constants satisfying*

- (a)  $\{a_n/n^{1/\gamma}\}$  is nondecreasing ultimately for some  $\gamma \in (0, 2)$ ;
- (b)  $\sup_n (a_{2n}/a_n) < \infty$ .

If  $\sum_n P(|Y_1| > a_n) < \infty$ , then

$$\lim_{n \rightarrow \infty} \frac{\sum_{j=1}^n Y_j - nEY_1 I(|Y_1| \leq a_n)}{a_n} = 0 \quad \text{a.s.}$$

If  $\sum_n P(|Y_1| > a_n) = \infty$ , then

$$\limsup_{n \rightarrow \infty} \frac{|\sum_{j=1}^n Y_j - c_n|}{a_n} = \infty \quad \text{a.s.}$$

for every sequence  $\{c_n\}$ .

Lemma 3.1 follows from Bingham et al. (1987, Theorem 1.5.6, p. 25), and Lemma 3.2 follows from Mori (1977, Theorem 1 and Lemmas 1 and 2).

The following lemma plays a central role in the proofs of our theorems.

**Lemma 3.3.** *Let  $\{p_n\}$  be a sequence of non-decreasing numbers with  $p_n \geq 1$  and  $\sup_n (p_{2n}/p_n) < \infty$ . Then, under (1.5), we have with probability one*

$$\limsup_{n \rightarrow \infty} \frac{|S_n - A_n|}{B(np_n)} = 0 \tag{3.1}$$

provided that  $\sum_{n=1}^\infty (1/np_n) < \infty$ .

**Proof.** Note that under (2.1),  $G(x) = P(|X| > x)$  is a regularly varying function with index  $-\alpha$  at infinity. From De Haan (1970) or Bingham et al. (1987),  $B(x)$  is a regularly varying function with

index  $1/\alpha$  at infinity, and thus  $B(x)$  has the following representation:

$$B(x) = c_1(x)x^{1/\alpha} \exp \left\{ \int_1^x \frac{b_1(u)}{u} du \right\},$$

where  $\lim_{x \rightarrow \infty} c_1(x) = c_1 \in (0, \infty)$  and  $\lim_{x \rightarrow \infty} b_1(x) = 0$ . Using the properties of regular variation (see Bingham et al., 1987), we have

$$\lim_{x \rightarrow \infty} xG(B(x)) = 1. \tag{3.2}$$

Define

$$g(x) = c_1 x^{1/\alpha} \exp \left\{ \int_1^x \frac{b_1(u)}{u} du \right\}.$$

It is easy to check that  $G(x) \sim g(x)$  as  $x \rightarrow \infty$  and that for any given  $\gamma \in (\alpha, 2)$ , there exists an  $x_\gamma > 0$  such that  $g(x)/x^{1/\gamma}$  is increasing in  $(x_\gamma, \infty)$ . Therefore  $\lim_{n \rightarrow \infty} (B(n)/g(n)) = 1$  and the sequence  $a_n = g(np_n)$  satisfies the conditions of Lemma 3.2. Moreover,

$$G(a_n) = G(g(np_n)) \sim G(B(np_n)) \sim \frac{1}{np_n}.$$

This implies that  $\sum_{n=1}^\infty P(|X| > a_n) < \infty$ . So, from Lemma 3.2,

$$\limsup_{n \rightarrow \infty} \frac{|S_n - nE(XI(|X| \leq a_n))|}{B(np_n)} = \limsup_{n \rightarrow \infty} \frac{|S_n - nE(XI(|X| \leq a_n))|}{a_n} = 0 \quad \text{a.s.} \tag{3.3}$$

Since  $p_n \rightarrow \infty$  and  $B(n) = o(B(np_n))$ , we have from (1.5) that

$$\frac{S_n - A_n}{B(np_n)} \xrightarrow{p} 0,$$

which, together with (3.3), yields

$$\frac{A_n - nE(XI(|X| \leq a_n))}{B(np_n)} \rightarrow 0.$$

Hence, the lemma follows from this approximation and (3.3).  $\square$

**Lemma 3.4.** *Let  $\{p_n\}$  be a sequence of non-decreasing numbers with  $p_n \geq 1$  and  $\sup_n (p_{2n}/p_n) < \infty$ . Suppose  $h \in BV[0, 1]$  with  $h(x_0) \neq 0$  for some  $x_0 \in (0, 1]$ , and  $h$  is continuous at  $x_0$ . Then under (1.5), with probability one*

$$\limsup_{n \rightarrow \infty} \frac{|S_n(h) - C_n|}{B(np_n)} = 0 \tag{3.4}$$

if  $\sum_{n=1}^\infty (np_n)^{-1} < \infty$ , and

$$\limsup_{n \rightarrow \infty} \frac{|S_n(h) - C_n|}{B(np_n)} = \infty \tag{3.5}$$

if  $\sum_{n=1}^\infty (np_n)^{-1} = \infty$ .

**Proof.** For any  $h \in BV[0, 1]$ , there exists a constant  $C$  such that  $|h(x)| \leq C$  and  $\sum_{k=1}^{n-1} |h(\frac{k}{n}) - h(\frac{k+1}{n})| \leq C$  for all  $n \geq 1$ . Therefore,

$$\begin{aligned} |S_n(h) - C_n| &= \left| \sum_{k=1}^n h\left(\frac{k}{n}\right) [X_k - (A_k - A_{k-1})] \right| \\ &= \left| \sum_{k=1}^{n-1} \left[ h\left(\frac{k}{n}\right) - h\left(\frac{k+1}{n}\right) \right] (S_k - A_k) + h(1)(S_n - A_n) \right| \\ &\leq 2C \max_{1 \leq k \leq n} |S_k - A_k|. \end{aligned}$$

Since (3.1) implies  $\limsup_{n \rightarrow \infty} [(\max_{1 \leq k \leq n} |S_k - A_k|)/B(np_n)] = 0$  due to the fact that  $B(np_n)$  is a non-decreasing function of  $n$ , (3.4) follows from Lemma 3.3 if  $\sum_{n=1}^{\infty} (np_n)^{-1} < \infty$ .

Next, we assume  $\sum_{n=1}^{\infty} (np_n)^{-1} = \infty$  and shall prove (3.5). It is easily seen that  $\limsup_{n \rightarrow \infty} [|S_n(h) - C_n|/B(np_n)]$  is measurable with respect to the tail- $\sigma$  field generated by the sequence  $\{X_n, n \geq 1\}$  (see Chow and Teicher, 1997, p. 64). From Kolmogorov zero-one law, there is a non-random constant  $d \in [0, \infty]$  such that with probability one

$$\limsup_{n \rightarrow \infty} \frac{|S_n(h) - C_n|}{B(np_n)} = d.$$

So we need to show  $d = \infty$ . Suppose at the moment that  $d < \infty$ . Then we will obtain some contradiction.

Since  $\sum_{n=1}^{\infty} (np_n)^{-1} = \infty$ , there exists a sequence of non-decreasing positive numbers,  $\{r_n\}$ , such that  $r_n \rightarrow \infty$  and

$$\sum_{n=1}^{\infty} \frac{1}{np_n r_n} = \infty. \tag{3.6}$$

Write  $B(x) = x^{1/\alpha}L(x)$ , where  $L(x)$  is a slowly varying function. By applying Lemma 3.1 with  $0 < \delta < 1/\alpha$  we obtain that  $B(np_n r_n)/B(np_n) \rightarrow \infty$  as  $n \rightarrow \infty$ . Then with probability one,

$$\lim_{n \rightarrow \infty} \frac{|S_n(h) - C_n|}{B(np_n r_n)} = 0.$$

Let  $\{X', X'_n, n \geq 1\}$  be an independent copy of  $\{X, X_n, n \geq 1\}$  and set  $X^s = X - X'$  and  $X_n^s = X_n - X'_n$  for  $n \geq 1$ . Then

$$\lim_{n \rightarrow \infty} \frac{|\sum_{k=1}^n h(\frac{k}{n})X'_k - C_n|}{B(np_n r_n)} = 0.$$

Further we have

$$\lim_{n \rightarrow \infty} \frac{|\sum_{k=1}^n h(\frac{k}{n})X_k^s|}{B(np_n r_n)} = 0. \tag{3.7}$$

Since the random variables  $X_n^s$  are symmetric, we could show that (3.7) implies

$$\sum_{n=1}^{\infty} P(|X^s| > B(np_n r_n)) < \infty \quad (3.8)$$

in a way similar to the proof of Theorem 2.1 of Chen (2002). Further, by using the following inequality:

$$P(|X - m(X)| > x) \leq 2P(|X^s| > x) \quad \text{for all } x > 0$$

where  $m(X)$  denotes the median of the random variable  $X$  (see Chow and Teicher, 1997, Lemma 10.1.1, p. 355), we obtain

$$\sum_{n=1}^{\infty} P(|X - m(X)| > B(np_n r_n)) < \infty. \quad (3.9)$$

Since

$$G(x + |m(X)|) \leq P(|X - m(X)| > x) \leq G(x - |m(X)|)$$

for any  $x > |m(X)|$ , and  $G(x)$  is regularly varying, it follows from (3.2) that as  $n \rightarrow \infty$

$$P(|X - m(X)| > B(np_n r_n)) \sim G(B(np_n r_n)) \sim \frac{1}{np_n r_n}.$$

So (3.9) implies  $\sum_{n=1}^{\infty} (np_n r_n)^{-1} < \infty$ , which is in contradiction with (3.6), i.e.,  $d = \infty$ . This completes the proof of this lemma.  $\square$

**Proof of Theorem 2.1.** Let “i.o.” denote “infinitely often”. It suffices to show that for every  $\varepsilon \in (0, 1)$

$$P \left( \left( \frac{|S_n(h) - C_n|}{B(\prod_{i=0}^{r-1} \lg_i(n))} \right)^{1/\lg_{r+1}(n)} > e^{(1+\varepsilon)/\alpha} \text{ i.o.} \right) = 0$$

and

$$P \left( \left( \frac{|S_n(h) - C_n|}{B(\prod_{i=0}^{r-1} \lg_i(n))} \right)^{1/\lg_{r+1}(n)} > e^{(1-\varepsilon)/\alpha} \text{ i.o.} \right) = 1,$$

or equivalently

$$P \left( |S_n(h) - C_n| > B \left( \prod_{i=0}^{r-1} \lg_i(n) \right) e^{((1+\varepsilon)/\alpha) \lg_{r+1}(n)} \text{ i.o.} \right) = 0 \quad (3.10)$$

and

$$P \left( |S_n(h) - C_n| > B \left( \prod_{i=0}^{r-1} \lg_i(n) \right) e^{((1-\varepsilon)/\alpha) \lg_{r+1}(n)} \text{ i.o.} \right) = 1. \quad (3.11)$$



Take  $p_n = [\lg_r(n)]^{\pm\varepsilon_1} \prod_{i=1}^r \lg_i(n)$  for  $\varepsilon_1 = \varepsilon/2$ , that is,  $np_n = [\lg_r(n)]^{1\pm\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n)$ . Since

$$\sum_{n=1}^{\infty} \frac{1}{[\lg_r(n)]^{1+\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n)} < \infty$$

and

$$\sum_{n=1}^{\infty} \frac{1}{[\lg_r(n)]^{1-\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n)} = \infty,$$

it is readily seen from Lemma 3.4 that

$$P \left( |S_n(h) - C_n| > B \left( [\lg_r(n)]^{1+\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n) \right) \text{ i.o.} \right) = 0$$

and

$$P \left( |S_n(h) - C_n| > B \left( [\lg_r(n)]^{1-\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n) \right) \text{ i.o.} \right) = 1.$$

Therefore, we only need to prove that for all large  $n$

$$B \left( [\lg_r(n)]^{1+\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n) \right) < B \left( \prod_{i=0}^{r-1} \lg_i(n) \right) e^{((1+\varepsilon)/\alpha)\lg_{r+1}(n)} \tag{3.12}$$

and

$$B \left( [\lg_r(n)]^{1-\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n) \right) > B \left( \prod_{i=0}^{r-1} \lg_i(n) \right) e^{((1-\varepsilon)/\alpha)\lg_{r+1}(n)}. \tag{3.13}$$

As in the proof of Lemma 3.4 write  $B(x) = x^{1/\alpha}L(x)$ , where  $L(x)$  is a slowly varying function. By applying Lemma 3.1 to function  $L(x)$  with  $d = \varepsilon_1/4$  we get for large  $n$

$$\begin{aligned} \frac{B([\lg_r(n)]^{1+\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n))}{B(\prod_{i=0}^{r-1} \lg_i(n))} &= [\lg_r(n)]^{(1+\varepsilon_1)/\alpha} \frac{L([\lg_r(n)]^{1+\varepsilon_1} \prod_{i=0}^{r-1} \lg_i(n))}{L(\prod_{i=0}^{r-1} \lg_i(n))} \\ &\leq 2[\lg_r(n)]^{(1+\varepsilon_1)/\alpha} [\lg_r(n)]^{d(1+\varepsilon_1)/\alpha} \\ &\leq 2[\lg_r(n)]^{(1+1.5\varepsilon_1)/\alpha} \\ &< [\lg_r(n)]^{(1+2\varepsilon_1)/\alpha} \\ &= e^{((1+\varepsilon)/\alpha)\lg_{r+1}(n)}, \end{aligned}$$

i.e., (3.12). Likewise, we can show (3.13). This completes the proof.  $\square$

**Proof of Theorem 2.2.** The proof follows from the lines of that of Theorem 2.1. The key is to show Lemma 3.4 holds for weighted sum  $\sum_{k=1}^n a_{n,k}X_k$ . In particular, if condition (C1) holds, then

$$\left| \sum_{k=1}^n a_{n,k}X_k - C_n \right| \leq C \max_{1 \leq k \leq n} |S_k - A_k|$$

for some constant  $C > 0$  (see the proof of Lemma 3.4), and thus (3.4) holds for  $\sum_{k=1}^n a_{n,k}X_k$ ; if condition (C2) holds, then  $\lim_{n \rightarrow \infty} [|\sum_{k=1}^n X_k^s|/B(np_n r_n)] = 0$ , an analogue of (3.7), yields the equation (3.8) which leads to some contradiction as in the proof of Lemma 3.4, that is, condition (C2) guarantees (3.5). The detailed proof is omitted.  $\square$

## References

- Bingham, N.H., Goldie, C.M., Teugels, J.L., 1987. *Regular Variation*. Cambridge University, New York.
- Chen, P., 2002. Limiting behavior of weighted sums with stable distributions. *Statist. Probab. Lett.* 60, 267–375.
- Chover, J., 1966. A law of the iterated logarithm for stable summands. *Proc. Amer. Math. Soc.* 17, 441–443.
- Chow, Y.S., Teicher, H., 1997. *Probability Theory: Independence, Interchangeability, Martingale*, 3rd Edition. Springer, New York.
- De Haan, L., 1970. On Regular Variation and Its Application to the Weak Convergence of Sample Extremes. *Math. Centre Tract* 32, Amsterdam.
- Gaposkin, V.F., 1965. The law of the iterated logarithm for Cesaro's and Abel's methods of summation. *Theory Probab. Appl.* 10, 420–431.
- Lai, T.L., Wei, C.Z., 1982. A law of the iterated logarithm for double arrays of independent random variables with application to regression and time series models. *Ann. Probab.* 10, 320–335.
- Li, D., Tomkins, R.J., 1996. Laws of the iterated logarithm for weighted sums of independent random variables. *Statist. Probab. Lett.* 27, 247–254.
- Loève, M., 1977. *Probability Theory I*, 4th Edition. Springer, New York.
- Mori, T., 1977. Stability for sums of i.i.d. random variables when extreme terms are excluded. *Z. Wahr. Geb.* 40, 159–167.
- Qi, Y., Cheng, P., 1996. A law of the iterated logarithm for partial sums in the domain of attraction of a stable law. *Chinese Ann. Math. Ser. A* 17, 195–206.
- Stadtmüller, U., 1984. A note on the law of the iterated logarithm for weighted sums of random variables. *Ann. Probab.* 12, 35–44.
- Tomkins, R.J., 1976. Strong limit theorems for certain arrays of random variables. *Ann. Probab.* 4, 444–452.